

Minimum Rates of Approximate Sufficient Statistics

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1 Sufficient Statistics, Motivation, and Main Contribution

Outline

- 1 Sufficient Statistics, Motivation, and Main Contribution
- 2 Problem Setup

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$$P_{X|\theta}(x) = \sum_{y \in \mathcal{Y}} P_{X|Y}(x|y)P_{Y|\theta}(y) = \sum_{y \in \mathcal{Y}} P_{X,Y|\theta}(x,y) \quad \forall(x, \theta)$$

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- In information theory language,

$$I(\theta; X) = I(\theta; f(X)) = I(\theta; Y).$$

Y provides as much information about θ as X does.

Examples

- $X^n = (X_1, \dots, X_n) \in \{0, 1\}^n$ is i.i.d. Bernoulli with parameter $\theta = \Pr[X_i = 1]$. Then

$$X^n \rightsquigarrow \frac{1}{n} \sum_{i=1}^n X_i \rightsquigarrow \theta$$

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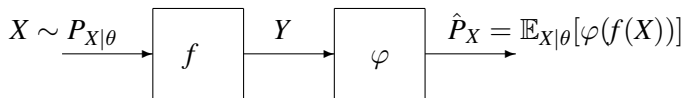
- Exponential family with natural parameter $\theta = (\theta_1, \dots, \theta_d)$

$$P_{X|\theta}^n(x^n) = P_X^n(x^n) \exp [\langle Y^{(n)}(x^n), \theta \rangle - nA(\theta)].$$

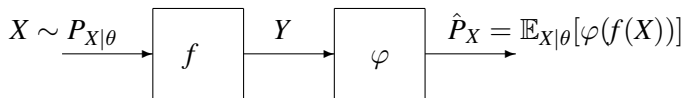
Vector of sufficient statistics $Y^{(n)}(x^n) = (Y_1^{(n)}(x^n), \dots, Y_d^{(n)}(x^n))$ with

$$Y_i^{(n)}(x^n) = \sum_{j=1}^n Y_i(x_j), \quad i = 1, \dots, d.$$

Another Interpretation : Exact Reproduction of $P_{X|\theta}$

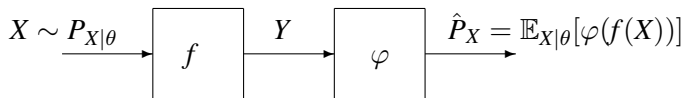


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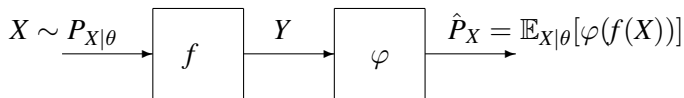
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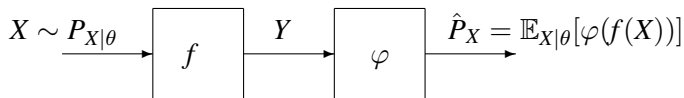


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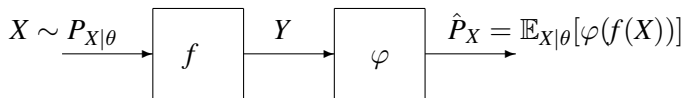
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$$\begin{aligned} \varphi \circ f \circ P_{X|\theta} &= \sum_{y \in \mathcal{Y}} (f \circ P_{X|\theta})(y) \varphi(y) \\ &= \sum_{y \in \mathcal{Y}} P_{X|\theta} \{x \in \mathcal{X} : f(x) = y\} P_{X|Y=y,\theta} = P_{X|\theta}. \end{aligned}$$

- Example 1: Binomial case. Since $\mathcal{X} = \{0, 1\}$, the sufficient statistic

$$\frac{1}{n} \sum_{j=1}^n X_j \in \left\{ \frac{0}{n}, \frac{1}{n}, \frac{2}{n}, \dots, \frac{n}{n} \right\}$$

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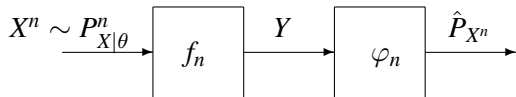
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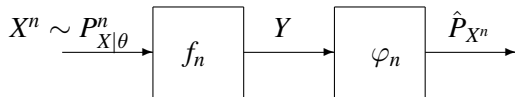
- Example 3: $\theta \in \Theta = [0, 1]$ is the unknown mean of a Gaussian. Sufficient statistics can take **uncountable** number of values.

Our Contribution

- Reduce d in n^d by relaxing exact recovery condition on $P_{X|\theta}^n$.



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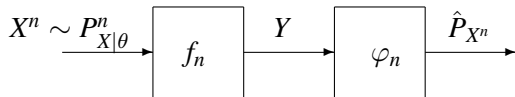
$$\varphi_n \circ f_n \circ P_{X|\theta}^n = P_{X|\theta}^n, \quad \forall n \in \mathbb{N},$$

we only require that

$$\overline{\lim}_{n \rightarrow \infty} \int_{\Theta} F\left(\underbrace{\varphi_n \circ f_n \circ P_{X|\theta}^n}_{\text{synthesized}}, \underbrace{P_{X|\theta}^n}_{\text{original}}\right) \mu(d\theta) \leq \delta, \quad \text{for some } \delta \geq 0,$$

where $F(\cdot, \cdot)$ is a “distance measure” between distributions.

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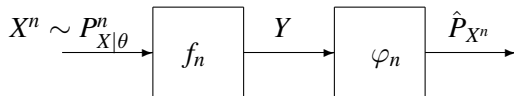
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- Often, we can reduce the exponent to $d/2$ and this is optimal.

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Definition of Code



Definition (Code)

A **size- M_n code** $\mathcal{C}_n = (f_n, \varphi_n)$ consists of

- A possibly stochastic encoder $f_n : \mathcal{X}^n \rightarrow \mathcal{Y}_n = \{1, \dots, M_n\}$;
- A decoder $\varphi_n : \mathcal{Y}_n \rightarrow \mathcal{P}(\mathcal{X}^n)$ (set of distributions on \mathcal{X}^n)

Definition of Error

Definition (Average Error)

The **average error** is a code $\mathcal{C}_n = (f_n, \varphi_n)$ is defined as

$$\begin{aligned}\varepsilon(\mathcal{C}_n) &:= \int_{\Theta} F\left(\varphi_n \circ f_n \circ P_{X|\theta}^n, P_{X|\theta}^n\right) \mu(d\theta) \\ &= \mathbb{E}_{\theta \sim \mu} \left[F\left(\varphi_n \circ f_n \circ P_{X|\theta}^n, P_{X|\theta}^n\right) \right]\end{aligned}$$

where $\mu(\cdot)$ is the prior distribution of θ .

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- Pinsker's inequality

$$\frac{\log e}{2} \|P - Q\|_1^2 \leq D(P\|Q)$$

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$$\lim_{n \rightarrow \infty} \frac{\log |\mathcal{C}_n|}{\log n} = r \iff |\mathcal{C}_n| \asymp n^r$$

Minimum Compression Rate: Properties

- Because $\delta \mapsto \mathbf{R}^{(i)}(\delta)$ is monotone

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- Typically for $\Theta \subset \mathbb{R}^d$,

$$R^{(i)}(\delta) = \frac{d}{2}.$$

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(v) Local asymptotic sufficiency of MLE

Theorem (Hayashi and Tan (2018))

- 1 Assume (i), (ii), (iv), and (v), under the variational distance criterion

$$\mathbf{R}^{(1)}(\delta) = \frac{d}{2} \quad \forall \delta \in [0, 2).$$

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$$\mathbf{R}^{(2)}(\delta) = \frac{d}{2} \quad \forall \delta \in \left[\frac{d}{2}, \infty\right).$$

- 3** *If in addition $\{P_{X|\theta}\}_{\theta \in \Theta}$ is an exponential family,*

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the memory requirement $d/2$ is asymptotically the same.

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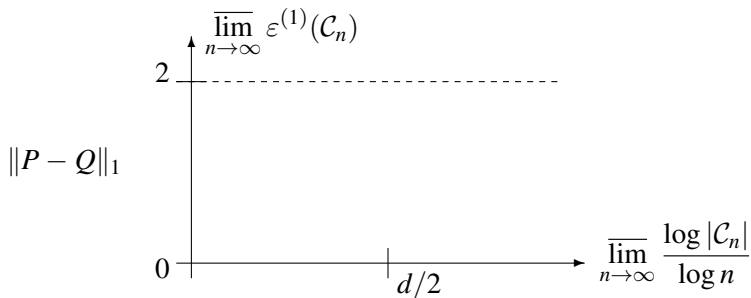
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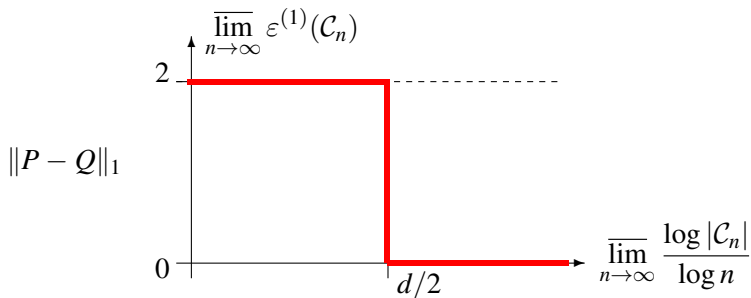
the memory requirement $d/2$ is asymptotically the same.

- This is known in information theory as a **strong converse**.

Main Result : Strong Converse



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Universal Coding, Information, Prediction, and Estimation

JORMA RISSANEN

Abstract—A connection between universal codes and the problems of prediction and statistical estimation is established. A known lower bound for the mean length of universal codes is sharpened and generalized, and optimum universal codes constructed. The bound is defined to give the information in strings relative to the considered class of processes. The earlier derived minimum description length criterion for estimation of parameters, including their number, is given a fundamental information theoretic justification by showing that its estimators achieve the information in the strings. It is also shown that one cannot do prediction in

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I. INTRODUCTION

THERE are three main problems in signal processing: prediction, data compression, and estimation. In the first, we are given a string of observed data points x_t , $t = 1, \dots, n$, one after another, and the objective is to predict for each t the next outcome x_{t+1} from what we have seen so far. In the data compression problem we are given a similar sequence of observations, each truncated to some finite precision, and the objective is to redescribe the data with a suitably designed code as efficiently as possible, i.e., with a short code length.

Manuscript received July 13, 1983; revised January 16, 1984. This work was presented in part at the IEEE International Symposium on Information Theory, St. Jovite, Canada, September 26–30, 1983.

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- Quantize the MLE similarly to Rissanen.

Weak Achievability for Relative Entropy: $\mathbb{R}^{(2)}\left(\frac{d}{2}\right) \leq \frac{d}{2}$

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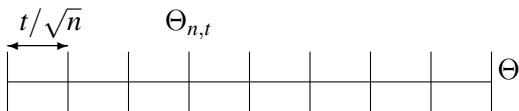
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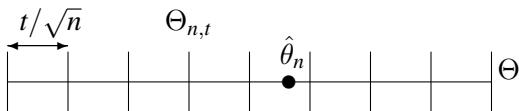
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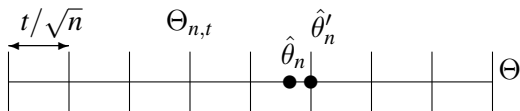
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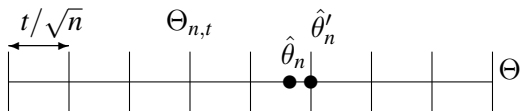
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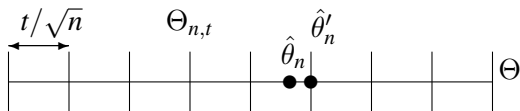
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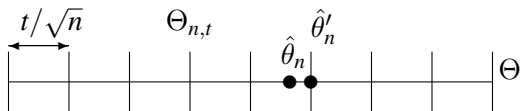
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$$\overline{\lim}_{n \rightarrow \infty} \varepsilon^{(2)}(\mathcal{C}_n) \leq \frac{d}{2}$$

by eventually taking $t \downarrow 0$. But error is non-vanishing. Weak achievability.

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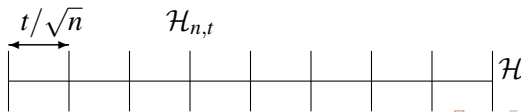
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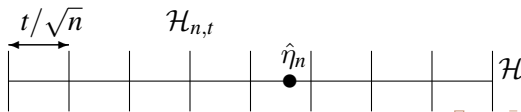
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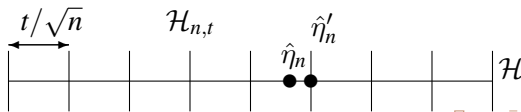
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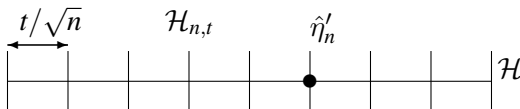
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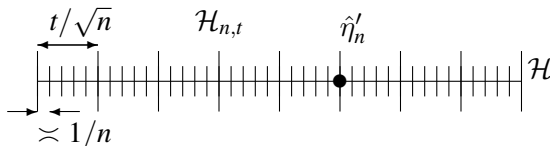
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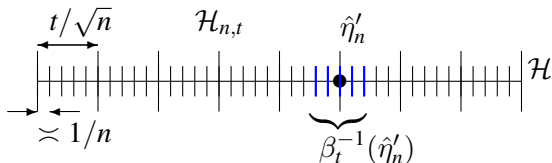
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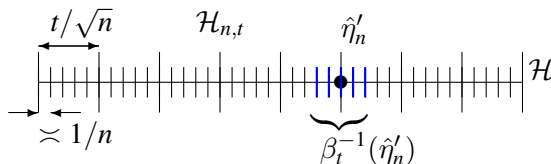
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- Asymptotic error under relative entropy is zero and $|\mathcal{H}_{n,t}| \asymp n^{d/2}$.

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Information-Theoretic Asymptotics of Bayes Methods

BERTRAND S. CLARKE AND ANDREW R. BARRON, MEMBER, IEEE

Abstract—In the absence of knowledge of the true density function, Bayesian models take the joint density function for a sequence of n random variables to be an average of densities with respect to a prior. We examine the relative entropy distance D_n between the true density and the Bayesian density and show that the asymptotic distance is $(d/2k \log n) + c$, where d is the dimension of the parameter vector. Therefore, the relative entropy rate D_n/n converges to zero at rate $(\log n)/n$. The constant c , which we explicitly identify, depends only on the prior density function and the Fisher information matrix evaluated at the true parameter value. Consequences are given for density estimation, universal data compression, composite hypothesis testing, and stock-market portfolio selection.

I. INTRODUCTION

THE RELATIVE entropy is a mathematical expres-

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The relative entropy rate between the true distribution and the mixture of distributions has been examined by Barron [4]. It is shown that if the prior assigns positive mass to the relative entropy neighborhoods $\{\theta: D(P_{\theta_n} \| P_{\theta_0}) < \epsilon\}$, $\epsilon > 0$, then

$$\lim_{n \rightarrow \infty} \frac{1}{n} D(P_{\theta_n} \| M_n) = 0 \quad (1.1)$$



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$$\varepsilon^{(1)}(\mathcal{C}_n) \rightarrow 0$$

and the uniform continuity of mutual information [Zhang (1997)]:

$$|I_P(A; B) - I_{P'}(A; B)| \leq 3\nu \log(|\mathcal{A}||\mathcal{B}| - 1) + 3H(\nu)$$

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$$\varepsilon^{(1)}(\mathcal{C}_n) \rightarrow 0$$

and the uniform continuity of mutual information [Zhang (1997)]:

$$\left| \underbrace{I_P(A; B)}_{\approx \frac{d}{2} \log n} - \underbrace{I_{P'}(A; B)}_{\leq H(Y_n) \leq \log |\mathcal{Y}_n|} \right| \leq \underbrace{3\nu \log(|\mathcal{A}||\mathcal{B}| - 1) + 3H(\nu)}_{\text{small}}$$

where

$$\nu = \frac{1}{2} \|P - P'\|_1.$$

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- Define $\mathcal{S} = \left\{ \theta \in \Theta : \left\| P_{X|\theta}^n - (\varphi \circ f)(\theta) \right\|_1 \leq 2 - \frac{\alpha}{2} \right\}$. Markov inequality says

$$\mu(\mathcal{S}) \geq \frac{\alpha}{4 - \alpha} > 0.$$

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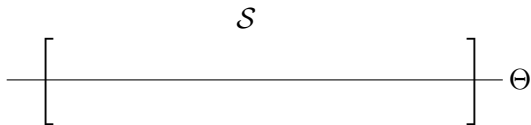
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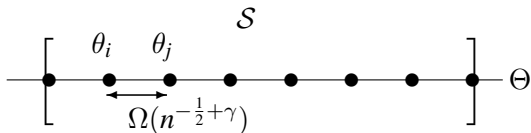
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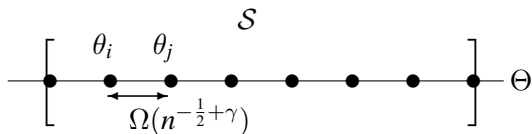
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- Because separation is $\Omega(n^{-\frac{1}{2} + \gamma})$, there exists **disjoint** $\mathcal{D}_i \subset \mathcal{X}^n$, $i = 1, \dots, \frac{5}{\alpha}M_n$ such that

$$P_{X|\theta_i}^n(\mathcal{D}_i) \geq 1 - \epsilon, \quad \text{for any } \epsilon \in (0, 1).$$

Follows by weak law of large numbers.

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- We have

$$(\varphi \circ f(\theta_i))(\mathcal{D}_i) \geq \frac{\alpha}{4} - \epsilon, \quad \forall i = 1, \dots, \frac{5}{\alpha} M_n.$$

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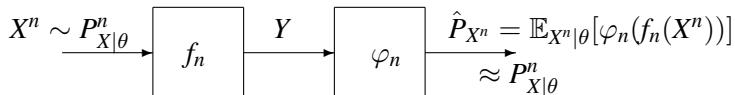
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Contradiction if $0 < \epsilon < \frac{\alpha}{20}$.

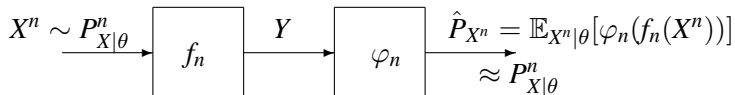
- 1 Sufficient Statistics, Motivation, and Main Contribution
- 2 Problem Setup
- 3 Main Result and Interpretation
- 4 Proof Ideas : Achievability
- 5 Proof Ideas : Converse (Impossibility)
- 6 Conclusion**

Concluding Remarks



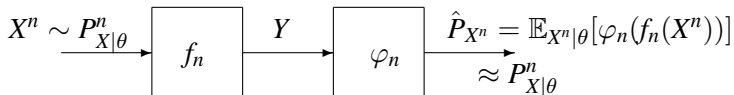
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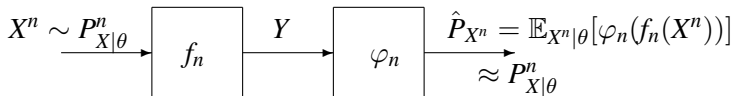
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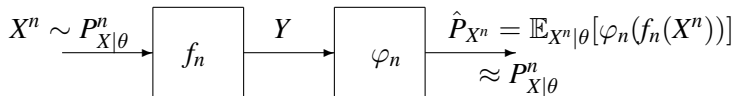
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